

Computational Approaches to Stochastic Portfolio Optimization Under Regime-Switching Dynamics

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Abstract

This dissertation develops computational frameworks for optimizing investment portfolios when financial markets exhibit abrupt shifts in volatility and correlation structure — a phenomenon known as regime switching. Traditional mean-variance optimization, rooted in Gaussian and stationary assumptions, systematically underestimates tail risk during turbulent regimes and over-allocates to assets whose correlations converge in downturns. We propose a class of hidden Markov regime-switching models coupled with stochastic gradient-based solvers that jointly estimate latent regimes and construct forward-looking efficient frontiers.

The core contribution is a two-stage procedure: first, a Bayesian changepoint detection algorithm segments historical returns into distinct volatility regimes using reversible-jump Markov chain Monte Carlo; second, a scenario-based convex optimizer — implemented via the alternating direction method of multipliers — solves the resulting non-Gaussian portfolio problem. We prove that the optimizer converges to the true conditional value-at-risk frontier under mild regularity conditions on the regime process.

Empirical evaluation uses synthetic data calibrated to the post-2008 regulatory environment and backtests on a 30-year panel of equity, fixed-income, and commodity indices. The regime-aware portfolios reduce maximum drawdown by 18–24% relative to static Markowitz benchmarks during stress periods while maintaining comparable Sharpe ratios in calm markets. We extend the framework to incorporate transaction costs, short-sale constraints, and ESG tilts, demonstrating its flexibility for institutional asset allocators. The methods are packaged as an open-source R library, `regimePort`, with bindings to `Stan` for posterior sampling and

CVXR for disciplined convex programming.

Keywords: portfolio optimization, regime switching, hidden Markov models, convex optimization, conditional value-at-risk, Bayesian changepoint detection, ADMM, risk management.

Dedication

*To my parents, who taught me that curiosity
is the only compass worth following.*

Table of Contents

Abstract	3
Dedication	5
List of Tables	3
List of Tables	3
List of Figures	4
List of Figures	4
1 Introduction	5
1.1 Motivation	5
1.2 The Regime-Switching Perspective	5
1.3 Contributions	6
1.4 Thesis Outline	7
2 Literature Review	8
2.1 Markowitz and Its Discontents	8
2.2 Hidden Markov Models in Finance	8
2.3 Convex Risk Measures	9
3 Bayesian Changepoint-Regularized HMM	10
3.1 Model Specification	10

3.2	Algorithm	11
3.3	Simulation Study	11
	Bibliography	13

List of Tables

2.1	Comparison of regime-switching estimation methods.	9
3.1	Reversible-jump MCMC for regime-switching HMM.	11
3.2	Regime count recovery across simulation settings ($N = 500$ replications). . .	12

List of Figures

Chapter 1

Introduction

1.1 Motivation

Modern portfolio theory, as formalized by Markowitz [10], remains the dominant paradigm for institutional asset allocation. Its elegance — distill risk into a covariance matrix, maximize expected return per unit of variance — has proven remarkably durable across seven decades of practice. Yet the 2008 financial crisis, the COVID-19 volatility spike of 2020, and the rate-hiking cycle of 2022–2023 each exposed the same structural vulnerability: correlations that appear modest in calm regimes collapse toward unity precisely when diversification is most needed.

This dissertation addresses a concrete question: can we build portfolio optimizers that are *regime-aware* — that anticipate shifts in the return-generating process rather than treating history as a single stationary draw? We answer affirmatively by combining tools from Bayesian time-series modeling and convex optimization into a unified computational pipeline.

1.2 The Regime-Switching Perspective

Financial returns exhibit well-documented stylized facts that violate the i.i.d. Gaussian benchmark: volatility clustering, fat tails, leverage effects, and time-varying correlations [4]. Hidden Markov models (HMMs) provide a parsimonious framework for capturing these

features. A K -state HMM posits that returns $\mathbf{r}_t \in \mathbb{R}^p$ are drawn from:

$$\mathbf{r}_t \mid s_t = k \sim \mathcal{N}_p(\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k), \quad k \in \{1, \dots, K\}, \quad (1.1)$$

where the latent state s_t evolves according to a first-order Markov chain with transition matrix $\mathbf{Q} = (q_{ij})$, and $q_{ij} = \Pr(s_t = j \mid s_{t-1} = i)$.

Definition 1 (Regime). A *regime* is a persistent state of the return-generating process characterized by a distinct pair $(\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$. Regimes typically correspond to macroeconomic environments: expansion (low volatility, positive drift), contraction (elevated volatility, negative drift), and crisis (extreme volatility, correlation breakdown).

The critical insight for portfolio construction is that the forward-looking distribution of returns is a *mixture*:

$$\Pr(\mathbf{r}_{T+1} \mid \mathcal{F}_T) = \sum_{k=1}^K \xi_T(k) \cdot \mathcal{N}_p(\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k), \quad (1.2)$$

where $\xi_T(k) = \Pr(s_{T+1} = k \mid \mathcal{F}_T)$ is the filtered state probability given information up to time T . When regimes are persistent, ξ_T concentrates mass on a single component, but the *possibility* of transition fattens the tails.

1.3 Contributions

This dissertation makes four principal contributions:

1. **Changepoint-regularized HMM estimation.** We introduce a reversible-jump MCMC sampler that jointly infers the number of regimes K and their parameters, penalizing superfluous states through a birth–death process on the transition matrix. This avoids the ad-hoc information-criterion selection common in the literature.
2. **Regime-conditional CVaR optimization.** We formulate the portfolio problem as

minimizing conditional value-at-risk under a mixture-of-Gaussians return model, transforming a non-convex objective into a second-order cone program solvable via ADMM. Convergence guarantees are established under α -mixing conditions on the regime chain.

3. **Empirical calibration and backtesting.** Using a synthetic data engine calibrated to post-2008 market structure and out-of-sample tests on a multi-asset panel spanning 1994–2024, we document consistent improvements in downside-risk metrics without sacrificing upside capture.
4. **Open-source software.** The `regimePort` package for R provides a modular implementation interfacing with `Stan` for Bayesian computation and `CVXR` for convex optimization, lowering the barrier for practitioners.

1.4 Thesis Outline

Chapter 2 surveys the literature on regime-switching models in finance, convex risk measures, and computational Bayesian methods. Chapter 3 develops the changepoint-regularized HMM estimator with full algorithmic details and convergence diagnostics. Chapter 4 formulates the regime-conditional CVaR portfolio problem and derives the ADMM algorithm. Chapter 5 presents the simulation study and empirical backtests. Chapter 6 discusses extensions to transaction costs, ESG constraints, and multi-period rebalancing. Chapter 7 concludes with limitations and directions for future work.

Chapter 2

Literature Review

2.1 Markowitz and Its Discontents

Markowitz [10] framed portfolio selection as a quadratic program: choose weights $\mathbf{w} \in \mathbb{R}^p$ to

$$\min_{\mathbf{w}} \mathbf{w}^\top \Sigma \mathbf{w} \quad \text{s.t.} \quad \mathbf{w}^\top \boldsymbol{\mu} \geq \mu_0, \quad \mathbf{w}^\top \mathbf{1} = 1. \quad (2.1)$$

The solution delivers the familiar efficient frontier — a parabola in mean-variance space. Its weaknesses are equally familiar: $\boldsymbol{\mu}$ and Σ are notoriously difficult to estimate (the “Markowitz error-maximization” problem [11]), and the Gaussian assumption ignores tail dependence.

Subsequent decades produced a rich literature on robust optimization [5], shrinkage estimation [9], and Black–Litterman blending [3]. These refinements improve *static* estimates but do not address the core problem: the data-generating process itself is non-stationary.

2.2 Hidden Markov Models in Finance

Hamilton [8] introduced regime-switching to economics, modeling GDP growth as alternating between expansion and recession states. The framework migrated to finance through applications in volatility modeling [14], interest rate dynamics [1], and tactical asset allocation [7].

Estimation is typically performed via the expectation-maximization (EM) algorithm or Bayesian MCMC. The EM approach is fast but provides only point estimates; MCMC

yields full posterior distributions but scales poorly with the number of assets p . A key open problem — which Chapter 3 addresses — is determining K , the number of regimes, without resorting to information criteria that penalize complexity in an ad-hoc manner.

Table 2.1: Comparison of regime-switching estimation methods.

Method	Selects K ?	Uncertainty	Scalability
EM + BIC	Yes (criterion)	Point estimate only	$\mathcal{O}(TKp^2)$
Hamilton filter	No	Filtered probabilities	$\mathcal{O}(TK^2)$
Gibbs sampler	No	Full posterior	$\mathcal{O}(TK^2p^2)$ per iteration
RJ-MCMC (this work)	Yes (inference)	Full posterior incl. K	$\mathcal{O}(TK_{\max}^2p^2)$ per iter.

2.3 Convex Risk Measures

Value-at-Risk (VaR) — the α -quantile of the loss distribution — became the regulatory standard through the Basel Accords. Its theoretical deficiencies are well documented: VaR is not a coherent risk measure [2] because it fails the sub-additivity axiom. Conditional Value-at-Risk (CVaR), defined as the expected loss exceeding VaR,

$$\text{CVaR}_\alpha(L) = \mathbb{E}[L \mid L \geq \text{VaR}_\alpha(L)], \quad (2.2)$$

is coherent and convex [12]. Rockafellar and Uryasev [13] showed that CVaR minimization can be cast as a linear program when the return distribution is represented by scenarios — an insight we exploit in Chapter 4, extending it to mixture-of-Gaussians representations.

Definition 2 (Coherent Risk Measure). A risk measure $\rho : \mathcal{X} \rightarrow \mathbb{R}$ is *coherent* if it satisfies (i) monotonicity: $X \leq Y \Rightarrow \rho(X) \geq \rho(Y)$, (ii) sub-additivity: $\rho(X + Y) \leq \rho(X) + \rho(Y)$, (iii) positive homogeneity: $\rho(\lambda X) = \lambda\rho(X)$ for $\lambda \geq 0$, and (iv) translational invariance: $\rho(X + c) = \rho(X) - c$.

Chapter 3

Bayesian Changepoint-Regularized HMM

3.1 Model Specification

We observe a p -dimensional return series $\{\mathbf{r}_t\}_{t=1}^T$. The generative model is:

Hierarchical Model

$$K \sim \text{Poisson}^+(\lambda_0) \quad (3.1)$$

$$q_{ii} \sim \text{Beta}(a_0, b_0), \quad i = 1, \dots, K \quad (3.2)$$

$$\boldsymbol{\mu}_k \sim \mathcal{N}_p(\mathbf{0}, \tau^{-1} \mathbf{I}), \quad k = 1, \dots, K \quad (3.3)$$

$$\boldsymbol{\Sigma}_k \sim \mathcal{W}_p^{-1}(\nu_0, \mathbf{S}_0), \quad k = 1, \dots, K \quad (3.4)$$

$$s_1 \sim \text{Categorical}(\boldsymbol{\pi}_0)$$

$$s_t \mid s_{t-1} \sim \text{Categorical}(\mathbf{q}_{s_{t-1}, \cdot}), \quad t = 2, \dots, T$$

$$\mathbf{r}_t \mid s_t \sim \mathcal{N}_p(\boldsymbol{\mu}_{s_t}, \boldsymbol{\Sigma}_{s_t}), \quad t = 1, \dots, T$$

The innovation is the prior on K : a zero-truncated Poisson with rate λ_0 , combined with a reversible-jump mechanism [6] that proposes birth, death, and split-merge moves on the state space.

3.2 Algorithm

Algorithm ?? sketches the main loop. At each iteration, with probabilities b_K , d_K , and $1 - b_K - d_K$, we propose a birth, death, or within-model move respectively.

Table 3.1: Reversible-jump MCMC for regime-switching HMM.

Step	Operation
1	Initialize $K^{(0)}$, $\theta^{(0)} = \{\mu_k, \Sigma_k, \mathbf{Q}\}$, and $s_{1:T}^{(0)}$
2	for $m = 1, \dots, M$ do
3	Sample $u \sim \text{Uniform}(0, 1)$
4	if $u < b_K$: propose birth move, accept with prob. α_{birth}
5	else if $u < b_K + d_K$: propose death move, accept with prob. α_{death}
6	else : update θ via Gibbs, sample $s_{1:T}$ via forward-filtering backward-sampling
7	end for

The birth move draws a new state ($K + 1$) by splitting an existing regime randomly; the death move merges two adjacent regimes. Acceptance probabilities are derived from the usual Metropolis–Hastings–Green ratio, with the Jacobian of the dimension-changing transformation absorbed into the proposal ratio [6].

3.3 Simulation Study

We evaluate the sampler on synthetic data with known ground truth. Table 3.2 reports recovery rates for K across 500 replications under three data-generating processes.

Table 3.2: Regime count recovery across simulation settings ($N = 500$ replications).

True K	$T = 500$	$T = 1000$	$T = 2000$	Avg. Runtime (s)
$K = 2$	0.94	0.97	0.99	42.3
$K = 3$	0.88	0.93	0.97	67.8
$K = 4$	0.79	0.87	0.94	98.2

Recovery is excellent for $K \leq 3$ with $T \geq 1000$, broadly adequate for $K = 4$, and the sampler correctly avoids over-fitting (the mean posterior K exceeds the truth by at most 0.18 across settings).

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